

ResearchSpace@Auckland

Version

This is the Accepted Manuscript version. This version is defined in the NISO recommended practice RP-8-2008 <u>http://www.niso.org/publications/rp/</u>

Suggested Reference

Taylor, W.J., Fransen, J., Jansen, T. L., Dalbeth, N., Schumacher, H. R., Brown, M., . . . Neogi, T. (2015). Study for Updated Gout Classification Criteria: Identification of Features to Classify Gout. *Arthritis Care and Research*, *67*(9), 1304-1315. doi: 10.1002/acr.22585

Copyright

Items in ResearchSpace are protected by copyright, with all rights reserved, unless otherwise indicated. Previously published items are made available in accordance with the copyright policy of the publisher.

This is the peer reviewed version of the following article: Taylor,W.J., Fransen, J., Jansen, T. L., Dalbeth, N., Schumacher, H. R., Brown, M., . . . Neogi, T. (2015). Study for Updated Gout Classification Criteria: Identification of Features to Classify Gout. *Arthritis Care and Research*, *67*(9), 1304-1315., which has been published in final form at <u>10.1002/acr.22585</u>

This article may be used for non-commercial purposes in accordance with <u>Wiley</u> <u>Terms and Conditions for Self-Archiving</u>.

http://www.sherpa.ac.uk/romeo/issn/2151-464X/

https://researchspace.auckland.ac.nz/docs/uoa-docs/rights.htm

Running head: Title: Authors: Institutions:

This article has been accepted for publication and undergone full peer review but has not been through the copyediting, typesetting, pagination and proofreading process which may lead to differences between this version and the Version of Record. Please cite this article as an 'Accepted Article', doi: 10.1002/acr.22585 © 2015 American College of Rheumatology Received: Nov 11, 2014; Revised: Mar 03, 2015; Accepted: Mar 10, 2015 This article is protected by copyright. All rights reserved.

Features that classify gout

Study for Updated Gout Classification Criteria (SUGAR): identification of features to classify gout

William J. Taylor PhD FRACP¹, Jaap Fransen PhD², Tim L. Jansen MD², Nicola Dalbeth MD FRACP³, H. Ralph Schumacher MD⁴, Melanie Brown MHealSc¹, Worawit Louthrenoo MD⁵, Janitzia Vazquez-Mellado MD PhD⁶, Maxim Eliseev MD⁷, Geraldine McCarthy MD FRCPI⁸, Lisa K. Stamp MD FRACP¹⁰, Fernando Perez-Ruiz MD¹¹, Francisca Sivera MD¹²,Hang-Korng Ea MD^{13 14}, Martijn Gerritsen MD¹⁵, Carlo Scire MD PhD¹⁶, Lorenzo Cavagna MD¹⁷, Chingtsai Lin MD¹⁸, Yin-Yi Chou MD¹⁹, Anne-Kathrin Tausche MD²⁰, Ana Beatriz Vargas-Santos MD²¹, Matthijs Janssen MD²², Jiunn-Horng Chen MD PhD^{23 24}, Ole Slot MD²⁵, Marco A. Cimmino MD²⁶, Till Uhlig MD²⁷, Tuhina Neogi MD PhD FRCPC²⁸

¹University of Otago Wellington, New Zealand.

²Radboud University Medical Centre, Nijmegen, Netherlands.

³University of Auckland, New Zealand.

⁴University of Pennsylvania and VA Medical Center,

Philadelphia PA.

⁵Division of Rheumatology, Department of Internal Medicine,

Faculty of Medicine, Chiang Mai University, Chiang Mai,

Thailand.

Acc

⁶Servicio de Reumatología, Hospital General de México, México City, México. ⁷Nasonova Research Institute of Rheumatology of Russia, Moscow, Russia. ⁸School of Medicine and Medical Science, University College Dublin, Ireland ⁹Mater Misericordiae University Hospital, Eccles St, Dublin, Ireland ¹⁰Department of Medicine, University of Otago Christchurch, PO Box 4345, Christchurch 8140, New Zealand ¹¹Rheumatology Division, Hospital Universitario Cruces & BioCruces Health Research Institute, Vizcaya, Spain. ¹²Department Reumatologia, Hospital General Universitario de Elda, Alicante, Spain. ¹³Univ. Paris Diderot, Sorbonne Paris Cité, UFR de Médecine, F-75205 Paris, France; INSERM, UMR 1132, Hôpital Lariboisière, F-75475 Paris, France ¹⁴Assistance Publique-Hôpitaux de Paris (AP-HP), Hôpital Lariboisière, Service de Rhumatologie, Centre Viggo Petersen, Pôle Appareil Locomoteur, 2, Rue Ambroise Paré, F-75010 Paris, France

¹⁵Amsterdam Rheumatology Immunology Center (ARC), Department of Rheumatology, Westfries Gasthuis, Hoorn, The Netherlands.

¹⁶Epidemiology Unit, Italian Society for Rheumatology (SIR), via Turati 40, 20121, Milan, Italy.

¹⁷Division of Rheumatology, University and IRCCS PoliclinicoS. Matteo Foundation, Viale Golgi 3, 27100, Pavia, Italy.

¹⁸Division of Rheumatology and Immunology, Taichung Tzu Chi Hospital, Buddhist Tzu Chi Medical Foundation, Taiwan, ROC.

¹⁹Taichung Veterans' General Hospital, 160 Taichung Road Section 3, Taichung, Taiwan 40705, ROC.

²⁰Division of Rheumatology, Department of Internal Medicine
III, University Hospital Carl Gustav Carus, Fetscherstrasse 74,
D-01307 Dresden, Germany.

²¹Division of Rheumatology, Department of Internal Medicine,Universidade do Estado do Rio de Janeiro, Brasil.

²²Rijnstate Hospital, Arnhem, The Netherlands.

²³School of Medicine, China Medical University, Taichung, Taiwan, ROC.

²⁴Department of Internal Medicine, China Medical University Hospital, Taichung, Taiwan, ROC

Corresponding author: Keywords: Acknowledgements:

²⁵Department of Rheumatology, Copenhagen University Hospital Glostrup DK-2600, Denmark. ²⁶Research Laboratory and Academic Division of Clinical Rheumatology, Department of Internal Medicine, University of Genova, Italy. ²⁷National Advisory Unit on Rehabilitation in Rheumatology, Department of Rheumatology, Diakonhjemmet Hospital, Box 23 Vinderen, 0319 Oslo, Norway. ²⁸Boston University School of Medicine, Boston MA. Associate Professor William Taylor, Department of Medicine, University of Otago Wellington, PO Box 7343, Wellington, New Zealand Phone +64 4 806 1801 Fax +64 4 389 5427 Email will.taylor@otago.ac.nz Gout, classification criteria, diagnostic ultrasound This study was supported by the American College of Rheumatology, European League against Rheumatism, Arthritis New Zealand, Association Rhumatisme et Travail, and Asociación de Reumatólogos del Hospital de Cruces. We gratefully acknowledge the help of Joung-Liang Lan,

Chien-Chung Huang, Po-Hao Huang, Hui-Ju Lin and Su-Ting

Chang (China Medical University Hospital, Taiwan), Anne

John Wiley & Sons, Inc.

Madigan (Dublin, Ireland), Yi-hsing Chen (Taichung, Taiwan), Alain Sanchez-Rodríguez and Eduardo Aranda-Arreola (Mexico City, Mexico), Viktoria Fana (Copenhagen, Denmark), Panomkorn Lhakum and Kanon Jatuworapruk (Chiang Mai, Thailand), Dianne Berendsen (Nijmegen, Netherlands), Femke Lamers-Karnebeek (Amsterdam, Netherlands), Olivier Peyr (Paris, France), Heidi Lunøe and Anne Katrine Kongtorp (Oslo, Norway), Geraldo da Rocha Castelar-Pinheiro (Rio de Janeiro, Brasil), Fatima Kudaeva (Moscow, Russia), Angelo Gaffo (Birmingham AL), Douglas White (Hamilton, New Zealand), Giovanni Cagnotto (Pavia, Italy) and Juris Lazovskis (Sydney, Canada) with data collection, crystal examination or patient referral. We are grateful to Eliseo Pascual (Alicante, Spain) for help with MSU observer certification and to Na Lu (Boston, MA) for help with the latent class analysis.

4265

Word count: Disclosures:

William Taylor – consultancy, honoraria or speaker fees from Pfizer (<\$10,000), educational grants from Abbvie, Roche, Pfizer (all <\$10,000); Lisa Stamp – consultancy, honoraria or speaker fees from AstraZeneca (<\$10,000); Fernando Perez-Ruiz – consultancy, honoraria or speaker fees from Menarini, Sobi, Pfizer, AstraZeneca, Novartis (all <\$10,000); H. Ralph Schumacher – consultancy, honoraria or speaker fees from Novartis, Regeneron, Metabolex and AstraZeneca (all<\$10,000); Martijn Gerritsen – speaking fees or consultancy

John Wiley & Sons, Inc.

Features that classify gout 6

from Sobi and Menarini (all<\$10,000); Anne-Kathrin Tausche – speaking fees for Berlin Chemie Menarini, expert testimony for Novartis, AstraZeneca, Ardea Bioscience (all <\$10,000); Ana Beatriz Vargas-Santos – speaking fees for AstraZeneca (<\$10,000); Tim Jansen – consultancy, honoraria or speaker fees from Abbvie, AstraZeneca, Bristol-Myers-Squib, Menarini and Roche; Marco Cimmino – consultancy, speaking fees or honoraria for Janssen Cilag (<\$10,000); Till Uhlig – honoraria for AstraZeneca and Sobi.

Acce

Abstract:

Objective: To determine which clinical, laboratory and imaging features most accurately distinguished gout from non-gout.

Methods: A cross-sectional study of consecutive rheumatology clinic patients with at least one swollen joint or subcutaneous tophus. Gout was defined by synovial fluid or tophus aspirate microscopy by certified examiners in all patients. The sample was randomly divided into a model development (2/3) and test sample (1/3). Univariate and multivariate association between clinical features and MSU-defined gout was determined using logistic regression modelling. Shrinkage of regression weights was performed to prevent over-fitting of the final model. Latent class analysis was conducted to identify patterns of joint involvement.

Results: In total, 983 patients were included. Gout was present in 509 (52%). In the development sample (n=653), these features were selected for the final model (multivariate OR) joint erythema (2.13), difficulty walking (7.34), time to maximal pain < 24 hours (1.32), resolution by 2 weeks (3.58), tophus (7.29), MTP1 ever involved (2.30), location of currently tender joints: Other foot/ankle (2.28), MTP1 (2.82), serum urate level > 6 mg/dl (0.36 mmol/l) (3.35), ultrasound double contour sign (7.23), Xray erosion or cyst (2.49). The final model performed adequately in the test set with no evidence of misfit, high discrimination and predictive ability. MTP1 involvement was the most common joint pattern (39.4%) in gout cases.

Conclusion: Ten key discriminating features have been identified for further evaluation for new gout classification criteria. Ultrasound findings and degree of uricemia add discriminating value, and will significantly contribute to more accurate classification criteria.

John Wiley & Sons, Inc.

Significance and innovation

Accurate classification criteria for gout that do not require MSU identification would be useful for clinical research.

Key clinical, imaging and laboratory features that distinguish between gout and nongout were identified in a cross-sectional, multi-national study.

The degree of uricemia and ultrasound findings contribute independently to the classification of gout and should be included in future classification criteria.

• Acce

Gout is the most common inflammatory arthritis in men and is increasing in prevalence (1, 2). Most gout worldwide is managed in primary care where disease identification seldom relies upon identification of monosodium urate (MSU) crystals because synovial fluid or polarizing light microscopy may not be easily obtainable (2). Therefore highly sensitive and specific classification criteria that do not require microscopic MSU crystal identification would be useful for clinical research conducted in epidemiological or primary care settings. Six classification criteria for gout have been developed, with the most widely used one being the 1977 American Rheumatism Association (ARA) criteria (3, 4).

However, there are significant limitations to the ARA criteria. These include the use of only patients with selected diagnoses (rheumatoid arthritis (RA), calcium pyrophosphate deposition disease (CPPD) and acute septic arthritis) as controls even though other rheumatic diseases may be amongst the group of patients or respondents for which such classification criteria for gout need to be applied. Only 47% of the controls had synovial fluid examination to confirm absence of MSU crystals.

The gold-standard chosen for the ARA classification criteria was physician diagnosis. In this way, the performance of synovial fluid microscopy could be examined. Only in 76 of 90 (84%) gout patients in whom synovial fluid was examined were MSU crystals identified. It is possible that excellent treatment of some patients led to clearance of MSU crystals, but the lower than expected sensitivity of synovial fluid analysis raises questions about the physician diagnosis (did the crystal-negative patients really have gout?) or the quality of the synovial fluid analysis. Many patients had incomplete data.

The ARA criteria were not tested in an external sample prior to publication, and thus were "preliminary" criteria. Subsequent external validation of the ARA survey criteria against a gold standard of synovial fluid analysis has been reported in two studies. In these studies the

John Wiley & Sons, Inc.

sensitivity was 70% and 80%, specificity was 78.8% and 64%, respectively (5, 6). Poor specificity is especially problematic in studies that seek to enroll patients into trials of new therapeutic agents, especially those with potential and unknown safety issues. It is important that patients accurately classified with gout are enrolled into studies to minimize inappropriate exposure to such drugs.

Due to the limitations of current criteria, ACR and EULAR have funded a project to update gout classification criteria (4). The <u>S</u>tudy for <u>U</u>pdated <u>G</u>out Cl<u>A</u>ssification C<u>R</u>iteria (SUGAR), undertaken as part of the gout classification project was designed to determine the performance of possible items that could discriminate between gout and non-gout, including diagnostic ultrasound findings. This study represents the first phase of the effort to develop updated classification criteria, aiming to identify the pertinent items to be considered for further evaluation and to assist with decision making in that evaluation. In addition, we examined the distribution of joint involvement amongst gout and non-gout patients to help distinguish patterns of joint involvement that should be considered in developing new criteria. It is important to re-iterate the difference between classification criteria, which are designed for entry into clinical studies, and diagnosis, which is a process that involves clinical care. The distinctions are more fully explained in existing literature (7-9). The work in the present study concerns classification criteria rather than diagnosis.

The second phase of the project consists of an expert-based selection and weighting process, informed by the SUGAR study results that will determine the new gout classification criteria that we aim to have endorsed by both ACR and EULAR. The second phase together with the final criteria will be reported separately.

Patients and Methods

Consecutive patients attending a rheumatology clinic with all of the following: a main complaint of joint(s) pain and swelling; GP or rheumatologist judges that gout is a possible differential diagnosis because of joint pain and swelling; patient has joint swelling currently or within the last 2 weeks or a possible tophus; there are no contraindications to arthrocentesis in the opinion of the clinical investigator, were enrolled into this crosssectional study. At the visit, clinical manifestations and a clinical diagnosis were recorded using standardized case record forms, prior to synovial fluid (or tissue) microscopy by a certified observer (see below for how observers were certified). Each centre received Ethics Committee Approval or Institutional Review Board approval according to local requirements. Informed consent was obtained from participants according to the requirements of local Ethics Committee or Institutional Review Board. The STAndards for the Reporting of Diagnostic accuracy studies (STARD) statement was used to guide reporting (10).

Gold-standard for gout (case definition)

All patients underwent arthrocentesis or tissue/tophus aspiration for polarizing microscopy to identify MSU crystals. All microscopic examination was undertaken by observers who had passed a 2-stage MSU-identification certification procedure, which consisted of a web-based erystal recognition test followed by examination of 5 to 8 vials of synovial fluid (SF) from the laboratories of Eliseo Pascual (European centres) or H. Ralph Schumacher (rest of the world) for those who passed the online test. The web-based test was strict and had a high non-pass rate (61%) (11). Each SF sample in the second stage needed to be correctly identified as demonstrating MSU crystals or not to achieve certification. SF samples with at least 10ml volume (to allow distribution of very small aliquots of the same specimens to all examinees) and typical findings for CPP crystals, MSU, and apatite (BCP, basic calcium phosphate) were selected. In addition, depot methylpredisolone crystals were added to a large osteoarthritic SF sample. Samples were pipetted into plastic tubes & sent with instructions to

John Wiley & Sons, Inc.

examine single drop specimens in the examinees microscope. Each SF sample was checked prior to shipping by express service to confirm stability of the crystals.

Cases of gout were defined as patients with MSU crystals identified by a certified observer. There were 42 certified observers involved in this study. Cases of non-gout were defined as patients without MSU crystals, irrespective of the clinical diagnosis. Synovial fluid/tissue microscopy by a certified observer was usually performed on the same day but could be done within 1 month of the visit to allow time to arrange an ultrasound guided aspiration. There was no restriction on the amount of fluid required from enrolled patients. Microbiological culture was sometimes not possible with small amounts of aspirated fluid, but this information was not mandatory. If arthrocentesis was not successful, investigators were able to repeat the procedure using US guidance. More than 1 joint could be aspirated according to the judgment of the clinical investigator, although this was uncommon. If no synovial fluid or tophus aspirate was obtained, the patient was not included in the study.

Potential classification items

Clinical data were collected at the index visit. Items were collected that included elements of existing classification criteria or had been ranked highly by physicians or patients in a previous Delphi exercise, which aimed specifically to identify potentially relevant diagnostic features of gout (12). Those items with a rating of 7 to 9 (agreed to be potentially discriminatory) from physicians or patients were selected.

All potential classification items (symptoms, signs, laboratory values and imaging results) were collected at the index visit and prior to synovial/tissue microscopy, using standardized case record forms. Currently swollen or tender joints were recorded on a homunculus.

Advanced imaging items (magnetic resonance imaging, computing tomography, dual energy computing tomography) were recorded if available but were not mandatory. Ultrasound and

John Wiley & Sons, Inc.

conventional radiography were intended to be mandatory but since data were usually collected within routine clinical care without specific funding, not every patient had radiography and ultrasound data available.

For participants with their first episode, items that referred solely to previous episodes were marked as non-applicable but predictors were constructed from items so as to include all episodes (current and previous).

Statistical analysis and sample size

The overall analytic plan is shown in supplementary Figure 1. Forty-five items were collected with the expectation of these reducing to 30 items for a multivariate analysis. It was planned that the dataset be randomly divided into development (2/3 of data) and test sets (1/3 of data), which was performed after all patients were included using a computerized random number generator. Using the rule of thumb for multivariate models of 10 participants per item divided by the smallest proportion of cases or controls (0.5) and 30 items for analysis, we estimated that 860 participants would be required for building a robust multivariate model (13).

Step 1: univariate analyses

In the development set, univariate logistic regression analysis was performed for each item, calculating the sensitivity (true positive rate, proportion of cases having the feature), 1-specificity (false positive rate, proportion of controls having the feature) and odds ratio (OR, odds of having the feature in cases, compared to controls) with its associated p-value.

After univariate analyses, 30 items were considered for inclusion in a multivariate model based on p-value (p<0.05; we did not use a more lenient p<0.1 at this stage due to the number of variables that were statistically significantly associated). Moreover, it was agreed that only items that were disease manifestations of gout would be included; thus risk indicators (as opposed to features of the disease itself) such as age, sex, diet and co-morbidities were not

John Wiley & Sons, Inc.

included. For items that were highly related (e.g., if there were >1 item regarding time to resolution or time to maximal pain), the item with the best clinical feasibility or, if equivalent, the item with the lowest p-value was selected.

Step 2: multivariate analyses

Two approaches were then taken to further develop the prediction model: firstly, a purely statistical approach starting with a full model containing all variables selected after univariate analyses, followed by backward selection using one-by-one deletion of items based on the size of the highest p-value of the regression coefficients until all p<0.2 (13). This was the statistically optimal model' used to gauge the performance of a clinical model. The second approach was a clinical hierarchical model using clinical reasoning and a forward selection procedure by collecting the major features of gout, sorted into domains of clinical information (joint inflammation, time course of symptoms, physical examination, laboratory test, and imaging). For that model, variables were selected from univariate analyses based on p-value (p < 0.05), OR (stronger the better), fit to the domains (while avoiding duplication), and feasibility with regards to assessment. Variables in a domain were included all at once and backward selection was applied until all p-value<0.20; then the variables of the next domain were added all at once and the procedure was repeated for that domain. Variables from previously added domains were only removed again if their p-value became >0.50. We chose a final model from these two approaches based on statistical parameters indicating better model fit (Akaike Information Criterion, AIC) and predictive accuracy (c-statistic) to determine the strength of associations between the included factors and presence of crystalproven gout.

Step 3: correcting for over-optimism

John Wiley & Sons, Inc.

We applied a standard method of shrinkage to reduce error in predicted values when the model is applied to new data and to avoid over-estimating the predictive ability of the model. The regression coefficients of the final model were reduced by multiplication with a shrinkage factor determined by performing a bootstrap procedure with 300 repetitions (14,

-15).

Step 4: testing model performance

The performance of the 'shrunken' model in the test set was analyzed using the c-statistic, Nagelkerke's R², the Hosmer-Lemeshow test and by plotting the predicted probabilities with the observed probabilities for groups classified according to deciles of the classification rule (calibration plot). The c-statistic is an estimate of discrimination, being equivalent to the area under the curve of a ROC curve, with values of 1 indicating perfect discrimination between gout and non-gout and 0.5 indicating discrimination no better than chance. The R² statistic is a measure of the difference between the observed class (gout/non-gout) and the class predicted by the prediction model (values of 0 indicating no explained variance and 1 indicating 100% of the variance in class membership are explained by the model). The Hosmer-Lemeshow test is a statistical test of model fit by assessing how well the observed class membership of subgroups correspond to class membership predicted by the model. A non-significant test indicates good model fit. The optimal cut point that corresponded to the inflexion point of the ROC curve signifying maximum sensitivity and specificity was also determined for this shrunken prediction model.

To further inform the decision-making process regarding which joints are most frequently involved during symptomatic episodes, patterns of joint involvement were analyzed in the development set of the SUGAR data. This was based upon currently swollen and tender joint involvement during the study visit as recorded on a homunculus following physician

John Wiley & Sons, Inc.

examination. First, we assessed the crude prevalence of each joint area with respect to reported tender vs. swollen joint involvement, and involvement of left-sided vs. right-sided joints in cases and controls separately. Once it was determined that prevalences were similar for both tender and swollen joints, and for left vs. right, tender and swollen joints were analyzed together (i.e., each joint was categorized as tender OR swollen), and both left-sided and right-sided joints were analyzed together in the same model.

We performed latent class analysis (LCA) on cases separately from the controls. LCA is a statistical approach that identifies underlying subgroups (i.e., latent classes or clusters) of individuals based upon their responses to a set of observed categorical variables. In this case, it was ascertainment of symptomatic involvement of each joint (yes vs. no) on a homunculus. This framework is the categorical analog to factor analysis, which is based on continuous variables.(16-18)

We used LCA to model the symptomatic joints to identify clusters representing a homogeneous subgroup that have a similar pattern of joint involvement. We first modeled all joints of the homunculus together, then repeated analyses limited to the following 4 joints areas: i) lower extremity joints only, ii) foot and ankle joints only, iii) upper extremity joints only, and iv) hand and wrist joints only. For each model, the optimal number of clusters was determined by statistical parameters (likelihood ration G^2 statistic, degrees of freedom, AIC and BIC). We obtained the prevalence of the identified clusters in each model, and the itemresponse probability of each item within each cluster.

Results

Twenty-five centres from 16 countries and 4 continents contributed data on 983 patients (509 cases, 472 non-cases), collected between January 2013 and April 2014. The majority, 702 (71.4%) were male. The mean (SD) age was 58.5 (17.2) years and median (interquartile

John Wiley & Sons, Inc.

range) duration of disease since first recalled symptoms was 4 (0.6 to 10) years (Table 1). In the development and test sets 30 (9%) and 17 (10%) of gout patients presented with their first episode of symptoms. Thirty-seven percent of the participants were non-Caucasian. Noncases had a clinical diagnosis of acute calcium pyrophosphate arthritis (109), spondyloarthritis (71), rheumatoid arthritis (70), osteoarthritis (69), undifferentiated arthritis (60), clinical gout (but MSU crystals were not observed by microscopy, 49), septic arthritis

(10), systemic lupus erythematosus (5), or other (31).

Missingness

MSU crystal identification was available for all patients. Clinical data collection and arthrocentesis was on the same day for 91% of patients and within 11 days for 99% of patients. Most variables had 0 or <1% missing only; serum uric acid level was missing for 8%, current location of tender joints, whether MTP1 was first ever joint involved were missing in <10%. Radiographs/ultrasound scans were not available from 12.4/18.5% of gout cases and 11.8/13.9% of non-gout cases. DECT and MRI were undertaken too infrequently to be usefully analyzed and were not considered further.

Univariate analysis

In the development sample (n=653), the univariate association of individual items with having MSU-proven gout is shown in Table 2. Overall, most items tested were associated with MSU-proven gout in univariate analyses, though some had weak associations in terms of magnitude of effect. The items with the highest OR were tophus on ultrasound (18.8), SUA >6mg/dl (0.36 mmol/l) (17.8), double contour sign on ultrasound (14.7), patient belief that he/she has gout (14.2), clinically evident tophus (10.0), dietary trigger (8.7), MTP1 ever involvement (8.1), radiographic erosion (6.7), male gender (6.4), joint erythema (6.1), marked tenderness (6.0), at least 2 episodes that start abruptly and resolve by 2 weeks (6.0),

John Wiley & Sons, Inc.

resolution of symptoms by 14 days (5.7), and difficulty walking or using joint (5.4). Additionally, we found a clear monotonically increasing relation of SUA levels with gout, with OR 5.9 for SUA 6 to 8mg/dl (0.36 to 0.48mmol/l) and 39.4 for SUA above 10mg/dl (0.6mmol/l). Nonetheless, a therapeutic cut-point of 6mg/dl (0.36mmol/l) was used in the final prediction model for this study to be more consistent with the dichotomous format of SUA in published criteria-sets and to be consistent with the data analysis that was presented to an expert panel in the second phase of the project.

Pattern of joint involvement

In the LCA model that included all currently involved (i.e., tender or swollen) joints among gout cases, 4 clusters emerged. The cluster with the highest prevalence was that of 1st MTP involvement, with a prevalence of 39.44% (Table 3). With just a slightly lower prevalence, the second most prevalent cluster was associated with predominantly knee or ankle/midfoot involvement (37.14%); the third was primarily elbow, wrist, and hand involvement (14.85%), and the fourth was polyarticular involvement, affecting virtually any joint (8.57%). This cluster pattern was in contrast to the non-gout patients, in which prevalence of knee involvement, predominantly monoarticular, appeared to be higher (56.83%), and no MTP pattern was seen (Table 3).

When limited to just lower extremity joints in the gout cases, monoarticular presentation of either knee, ankle, or 1st MTP was high (83.49%); limited to just the foot and ankle, 1st MTP monoarticular presentation was common (84.16%). For non-gout patients, when limited to the lower extremity, monoarticular involvement of the knee was most common (60.84%); when limited to the foot/ankle, 1st MTP monoarticular presentation was uncommon (6.36%). Upper extremity as well as hand/wrist clusters were similar for cases and controls.

Multivariate models

John Wiley & Sons, Inc.

The multivariate model developed from backward removal of items based solely on statistical criteria is shown in Table 4 along with the hierarchical models that included clinical reasoning. Two items were removed again from the hierarchical models because of large p-values (p>0.50). The final hierarchical model parameters are shown in Table 5. This model was chosen over the statistically derived model because of smaller AIC (303 vs. 677) and larger c-statistic (0.93 vs. 0.91).

We applied a shrinkage factor of 0.76 that was estimated using bootstrapping to the regression coefficients of the hierarchical model to form a final prediction model that provided insights into the strength of association between the included items and presence of MSU-proven gout (the regression coefficients or weights are shown in Table 5).

In the multivariate analyses, the items regarding marked tenderness and having at least 2 episodes that start abruptly and resolve within 2 weeks were no longer significantly associated with presence of crystal-proven gout. There was conceptual similarity between the item "having at least 2 episodes that start abruptly and resolve within 2 weeks" and 2 other variables that remained in the model, that of "maximal pain within 24 hours" and "resolution by 14 days". This redundancy led to removal of the first of these 3 items. Similarly, marked tenderness was likely analogous in concept to experiencing great difficulty in walking or inability to use affected joint. Further, while all ultrasound features were initially included in the hierarchical model, subsequently only DCS remained significant. Similarly, the only x-ray features that were retained for the multivariate model were erosions with selerotic margins and overhanging cortical edges and/or subcortical cysts.

In the test data-set (n=329, 50.5% with gout), a logistic regression model in which gout/nongout (based upon MSU presence) was the dependent variable and the model-derived 'shrunken' score (based upon the 'shrunken' regression coefficients derived in the

John Wiley & Sons, Inc.

development data-set; Table 5) was the independent variable, a close association between the score and the presence of gout was noted with R² 0.58, c-statistic 0.90 and Hosmer-Lemeshow Chi-square 7.11 (df=8), p=0.53. The model parameters for the version without imaging were R² 0.52, c-statistic 0.87 and Hosmer-Lemeshow Chi-square 1.90 (df=9), p=0.99. Using a statistically derived cut-point from the data-derived prediction model in which the sum of sensitivity and specificity was maximized, we found the sensitivity of the model was 83.9% (88.1% for the model without imaging) and specificity was 81.2% (71.6% for the model without imaging).

Discussion

In this study, we found that a large number of clinical and imaging parameters are highly associated with presence of MSU-crystal gout compared to other diseases associated with joint swelling. However, in multivariate analyses, some were no longer significantly associated with gout, while 10 factors remained independently predictive of the presence of gout, representing clinical, laboratory, and imaging features. Even so, the data-derived prediction model possessed sensitivity and specificity that were both less than 90%.

We also confirmed that MTP1 involvement was typical of gout, although the latent class analysis showed that knee/ankle/midfoot involvement was almost as common. Knee monoarthritis was more typical of non-gout, although also occurred frequently in gout cases. It should be acknowledged that this analysis was performed on currently involved joints, which does not necessarily capture the full spectrum of joint involvement that could 'ever' occur in a given patient.

In this study, we chose a therapeutic value of 6mg/dl (0.36mmol/l) as a cutoff to define high serum urate, but the association of serum urate with gout/non-gout varies continuously. This has been well described at a population level (19) but not previously in a diagnostic context.

John Wiley & Sons, Inc.

In the SUGAR data, we found that the odds ratio of having gout vs. non-gout increases nonlinearly with quartiles of the highest ever recorded serum urate from nearly 6 for the lowest quartile (less than 6 mg/dl, 0.36mmol/l) to 39 for the highest quartile (more than 10 mg/dl, 0.6mmol/l). This is shown graphically in Figure 1. Future proposals for gout classification criteria should possibly incorporate levels of serum urate rather than simple presence vs. absence of a high serum urate level.

This study took several steps to minimize bias, including a requirement that all subjects underwent a rigorous gold-standard procedure, collection of pre-defined items that were previously included in existing criteria sets and additional ones identified through Delphi methods, application of 'shrinkage' to estimated regression coefficients to minimize overfitting, and testing the derived criteria in a randomly selected subset of the data that did not contribute towards the development of the criteria. The sample size was large enough to test a range of possible items including key ultrasound features. Further, the subjects comprising our control group included a broader, relevant spectrum of diseases than prior classification criteria studies.

The main limitations to the interpretation of the results arise from the population that was studied. This was not a primary care population and all patients needed to have synovial fluid or tissue aspiration in order to be included in the study. This confers unavoidable spectrum bias, whereby the studied population likely had more severe than would be seen in primary care, and does not necessarily represent the full spectrum of clinical gout. In addition, this sample is biased towards chronic disease by inclusion of people with more persistent or recurrent features. Patients with persistent disease have more opportunity to accumulate disease features. Therefore the results of the study may not be generalizable to primary care settings.

John Wiley & Sons, Inc.

It is possible that patients with large joint disease were preferentially selected because of the need for joint aspiration. However, this is unlikely to have been a major factor since 74% of gout cases had MTP1 involvement at some time in their disease course, which is within the range of what is observed in other gout cohorts (20) and furthermore 128/1004 (13%) joint aspirations were from the MTP1 joint. There were 34% of gout cases with currently tender MTP1.

Whether these same items are also predictive of presence of crystal-proven gout in a sample that may have milder disease that is not tophaceous or disease that involves only small joints is presently not known. This is one of the reasons that additional work has been planned to develop the new gout classification criteria. We also could not test the performance of presence of MSU crystals itself given that our case-control status was defined by their presence (i.e., gold standard).

We emphasize that the prediction model described in this study was intended to inform the next stages of classification criteria development and are not the final ACR-EULAR endorsed gout classification criteria. The study has demonstrated which items and which combinations of items are most strongly associated with gout. In particular, it has demonstrated that ultrasound features, even when unstandardized and performed according to local practice, can significantly contribute to classification of gout.

Nonetheless, in order to overcome the limitation of spectrum bias, the need for all patients to have undergone the diagnostic gold-standard procedure and to allow expert opinion to be integrated into new classification criteria, the second phase of the project incorporating a paper patient exercise and an expert clinician workshop was undertaken (reported separately). Despite efforts to incorporate key diagnostic features of gout, classification of patients without searching for MSU crystals remains inaccurate with combined sensitivity and

John Wiley & Sons, Inc.

specificity of the prediction model derived from these data both under 90%. This highlights the importance of synovial fluid examination for diagnostic purposes in ordinary clinical care and cautions against the use of classification criteria for diagnosis. In situations where high positive predictive value is required, a high suspicion of gout (pre-test probability), criteria with lower sensitivity or a crystal diagnosis should be determined. Ultrasound features also contribute strongly to classification of gout and may have sufficient specificity to be useful in such situations.

Acc

John Wiley & Sons, Inc.

Acce

Table 1. Participant characteristics

		Gout (MSU positive, n=509)	Non-gout (MSU negative, n=474)
Time sin (years), 1	ce first episode of symptoms median (IQR)	6 (2 to 13)	3 (0.23 to 8)
Male (%)	440 (86%)	262 (55%)
Age, me	an (SD)	60 (15)	59 (16)
Clinical	Gout	494	49
diagnosi	s* Calcium pyrophosphate deposition disease	4	109
	Spondyloarthritis	1	71
	Osteoarthritis	3	69
	Rheumatoid arthritis	2	70
	Undifferentiated arthritis	1	33
	Septic arthritis	2	10
	Systemic Lupus Erythematosus	0	5
	Other	2	31
Ethnicity	White/European/Caucasian	347	276
-	African/Black	12	11
	Hispanic	23	21
	South Asian	47	38
	East Asian	67	113
	Pacific Island	3	1
	Other indigenous	3	4
	Other	7	10

* Clinical diagnosis was independent of MSU crystal identification

MSU = monosodium urate; IQR = interquartile range; SD = standard deviation

John Wiley & Sons, Inc.

Table 2. Univariate association of each variable with classification as gout in the developmentdata-set (n=653). Values are percentages for categorical items or mean (SD) for continuousvariables.

Item Description		Present in gout (N=346), %	Present in non-gout (n=307), %	OR	p-value	
Gender (male)		89%	55%	6.4	< 0.0001	
	0-4 hours	23%	19%	1.9		
Time to may pain (asymptoticada)	5-12 hours	28%	16%	2.0	<0.0001	
Time to max. pain (current episode)	13-24 hours	16%	12%	2.6	<0.0001	
	>24 hours	34%	52%	ref.		
	0-4 hours	31%	33%	1.9		
Time (5-12 hours	31%	14%	3.3	<0.0001	
Time to max. pain (previous episodes)	13-24 hours	16%	10%	4.5	<0.0001	
	>24 hours	22%	44%	ref.		
Pain level 0 to 10 (current episode)	- 1	8 (7-10)	8 (6-9)	1.1	0.0085	
Pain level 0 to 10 (current episode) 7 o	r higher	79%	70%	1.6	0.015	
Pain level 0 to 10 (previous episodes)	0	8 (7-10)	7 (4-9)	1.2	< 0.0001	
Pain level 0 to 10 (previous episodes) 7	7 or higher	76%	56%	2.6	< 0.0001	
At least 2 episodes that start abruptly a	nd conclude by				.0.0001	
2 weeks	5	83%	44%	6.0	<0.0001	
	0-24 hours	7%	6%	6.1		
	1-7 days	56%	28%	10.3	<0.0001	
	8-14 days	19%	11%	8.7		
Time for episodes to resolve	More than 14					
	days	11%	19%	2.6		
	Never					
	resolved	7%	37%	ref.		
Time for episodes to resolve	More than 14				<0.0001	
	days or never	18%	56%	0.17	-0.0001	
	1	9%	22%	ref.		
Number of episodes	2 to 5	24%	28%	3.1	< 0.0001	
	more than 5	66%	50%	2.0		
	yes	33%	12%	3.4	< 0.0001	
Stereotypical progression*	no	61%	80%	ref.		
	uncertain	6%	8%	1.0		
Complete resolution of at least one epis	sode	85%	56%	4.6	< 0.0001	
Hypertension or other CVS diseases		68%	45%	2.6	< 0.0001	
Tender joint count		1 (1-3)	1 (1-2)	1.04	0.06	
	1	53%	61%	0.59		
Tender joint count	2-4	30%	28%	0.74	0.0066	
	>4	17%	12%	ref.		
Swollen joint count		1 (1-3)	1 (1-2)	1.06	0.026	
Swollen joint count	1	53%	65%	0.38	0.0008	

John Wiley & Sons, Inc.

-

		2-4	32%	27%	0.55	
		> 4	15%	7%	ref.	
	Involved joints take longer to heal		25%	29%	0.84	0.33
	Patient belief he/she has gout		81%	23%	14.2	< 0.0001
	Told by doctor he/she has gout		84%	27%	14.3	< 0.0001
	Xray - erosion		42%	10%	6.7	< 0.0001
	Xray - asymmetric joint swelling		43%	24%	2.4	< 0.0001
	Xray - joint cyst without erosion		44%	17%	3.7	< 0.0001
	Xray – erosion and/or joint cyst		59%	21%	5.3	< 0.0001
	Any of these xray features present		66%	38%	4.7	< 0.0001
	US - double contour sign		58%	9%	14.7	< 0.0001
	US - tophus		45%	4%	18.8	< 0.0001
	US - snowstorm effusion		31%	9%	4.4	< 0.0001
	Any of these 3 US features present		75%	16%	15.3	< 0.0001
		grade 0	32%	42%	3.0	
	Dower Donnlor (US)	grade 1	25%	23%	2.1	0.0007
	Power Doppier (US)	grade 2	28%	29%	2.4	0.0097
		grade 3	15%	6%	ref	
	Patient taking ULT at time of SUA level		34%	8%	5.6	< 0.0001
	Tophus present or previously present		34%	5%	10.0	< 0.0001
' (MTP1 involvement was first ever sympto	m	54%	16%	6.0	< 0.0001
	MTP1 involvement ever occurred		74%	27%	7.6	< 0.0001
	MTP1 occurred in isolation ever		72%	24%	8.1	< 0.0001
	Maximal pain greater than 7		93%	81%	3.3	< 0.0001
	At least 1 episode resolved by 14 days		82%	44%	5.7	< 0.0001
	At least 1 episode involved 1-5 joints		72%	69%	1.2	0.37
	At least 1 episode involved joint erythem	a	87%	53%	6.1	< 0.0001
	At least 1 episode involved marked tende	rness to touch	92%	65%	6.0	< 0.0001
	At least 1 episode involved only 1 joint		95%	86%	3.2	0.0001
	At least 1 episode involved tarsal joints		57%	30%	3.0	< 0.0001
	Patient has history of kidney stone		15%	9%	1.8	0.017
		Yes	40%	10%	ref.	
	At least 1 episode responded to	No	13%	10%	0.34	<0.0001
	colchicine	Never used				<0.0001
		colchicine	47%	79%	0.15	
	At least 1 episode involved difficulty wal	king	97%	84%	5.4	< 0.0001
	At least 1 episode it was hard to sleep		90%	69%	4.2	< 0.0001
	At least 1 episode pain present at rest		91%	78%	3.0	< 0.0001
	At least 1 episode joint was throbbing or	burning	90%	72%	3.7	< 0.0001
	Disabled OR sleep OR rest pain OR throb	obing	99%	90%	9.4	< 0.0001
	At least 1 episode joint was hot to touch		95%	80%	4.8	< 0.0001
	At least 1 episode triggered by dietary factor		53%	11%	8.7	< 0.0001
	At least 1 episode there was joint swelling	g	97%	97%	1.1	0.79
	Current SUA (mmol/l)		0.47 (0.13)	0.34 (0.13)		

John Wiley & Sons, Inc.

Highest recorded SUA (m	mol/l)	0.56 (0.14)	0.38 (0.14)			
Highest SUA without ULT (mmol/l)		0.57 (0.13)	0.38 (0.13)			
Current SUA > 0.36 mmo	Current SUA > 0.36 mmol/l			5.5	< 0.00	
Highest recorded SUA > (0.36 mmol/l	93%	51%	12.6	< 0.00	
Highest SUA without UL	$\Gamma > 0.36 \text{ mmol/l}$	95%	51%	17.8	< 0.00	
	<0.36 mmol/l	7%	51%	Ref		
	0.36 to <0.48 mmol/l	19%	26%	5.9	-0.0001	
Hignest SUA category	0.48 to <0.60 mmol/l	38%	16%	18.3	<0.00	
	≥0.60 mmol/l	36%	7%	39.4	-	
Currently tender any MTF	21	33%	9%	4.8	< 0.00	
Currently tender any big t	oe IP	4%	0%	12.2	0.01	
Currently tender any joint	lesser toes	12%	7%	1.8	0.03	
Currently tender any midf	oot joint	8%	5%	1.8	0.07	
Currently tender any ankle	e joint	30%	17%	2.1	0.00	
Currently tender any knee	joint	42%	67%	0.27	< 0.00	
Currently tender any hip j	oint	1%	3%	0.18	0.02	
Currently tender any shou	lder joint	3%	6%	0.44	0.03	
Currently tender any elboy	w joint	11%	7%	1.7	0.05	
Currently tender any wrist	t ioint	15%	14%	1.0	0.8	
Currently tender any hand	ioint	20%	10%	2.1	0.00	
Currently tender any CMC		1%	1%	1.8	0.4	
	Any joint proximal to ankle	36%	68%	ref.		
Currently tender joint	Ankle or foot not MTP1	28%	20%	2.6	< 0.000	
location	MTP1	36%	12%	5.9		
		30%	10%	4.0	< 0.00	
		3%	0%	10.3	0.02	
		40%	48%	3.7	< 0.00	
		7%	4%	1.6	0.1	
		31%	17%	2.1	< 0.00	
As above except these are	for currently swollen joints	43%	67%	0.38	< 0.00	
	5	1%	4%	0.27	0.02	
		11%	5%	2.3	0.00	
		14%	14%	0.99	0.9	
		20%	8%	2.8	< 0.00	
		1%	1%	1.8	0.4	
	Any joint proximal to ankle	38%	71%	ref.		
Currently swollen joint	Ankle or foot not MTP1	30%	17%	3.2	< 0.000	
location			100/	5.0		

SD = standard deviation; OR = odds ratio; US = ultrasound; ULT = urate lowering therapy; MTP1 = first metatarsophalangeal joint; SUA = serum urate; CVS = cardiovascular; IP = interphalangeal joint; CMC = thumb carpometacarpal joint.

Table 3. Prevalence of joint involvement clusters among gout and non-gout

	Gout	Non-ge	out
Joint Pattern	Prevalence	Joint Pattern	Prevalence
MTP1	39.4%	Knee only	56.8%
Knee/ankle	37.1%	Any lower extremity	30.1%
		joint	
Elbow/wrist/hand	14.9%	Wrist/hand + knee	8.0%
Polyarticular	8.6%	Polyarticular	5.1%

MTP1 = first metatarsophalangeal joint

Acce

John Wiley & Sons, Inc.

Table 4. Multivariate model development. Two approaches are shown – by statistical selection only or by progressive addition of clinically informed groups of variables (see text for details).

Model name	Variables	AIC*	c-statistic (AUC)†
STATISTICAL SELECTION ONLY	Sex, onset of maximal pain, complete resolution of episode, xray features [‡] , US DCS, tophus, MTP1 involved, disability, hyperuricemia, location of current joint involvement	380	0.91
HIERARCHICAL MODEL DEVELOPMENT			
JOINT INFLAMMATION	Joint erythema, marked tenderness, joint warm to touch, disability	746	0.75
+ COURSE	+ onset of maximal pain with 24 hours, attack resolves within 14 days, at least 2 episodes	669	0.82
+ GOUT FEATURES	+ tophus, MTP1 ever involved, current tender joint location	531	0.89
+ URATE	+ highest recorded SUA>6mg/dl (0.36mmol/l), at least 2 episodes§	463	0.89
+ ULTRASOUND	+ US features	354	0.92
+ US only DCS	+ US DCS sign	355	0.92
+ XRAY	+ xray features	396	0.90
+ XRAY one sign	+ xray erosion or cyst, marked tenderness§	392	0.90
+ US AND XRAY	+ US DCS, xray erosion or cyst, marked tenderness§	303	0.93

*AIC, Akaike Information Criterion is an index of model fit adjusted for parsimony (lower values indicate better fit); † Area under the curve of a receiver operating characteristic curve (values of 0.5 indicate prediction of class membership no better than chance, values of 1.0 indicate perfect prediction of class membership); ‡ Xray features were erosion, subcortical cyst or asymmetric soft tissue swelling within a joint; §Item shown as deleted because it was removed due to p=0.47 (at least 2 episodes), p=0.22 (marked tenderness).

US DCS = ultrasound double-contour sign; MTP1 = first metatarsophalangeal joint; SUA = serum urate

Table 5. Item parameters of final model from development data-set (n=654)†

Variable	OR (95% CI)	p-value	Regression coefficient	Regression coefficient with shrinkage‡
Intercept			-5.70	-4.05
Joint erythema	2.13 (1.06, 4.29)	0.03	0.76	0.54
At least 1 episode involved difficulty walking	7.34 (1.17, 46.06)	0.03	2.00	1.42
Time to maximal pain less than 24 hours	1.32 (0.71, 2.47)	0.38	0.28	0.20
Resolution by 2 weeks	3.58 (1.85, 6.95)	0.0002	1.28	0.91
Tophus	7.29 (2.42, 21.99)	0.0004	1.99	1.41
MTP1 ever involved	2.30 (1.18, 4.49)	0.01	0.83	0.59
Location of currently	Other foot/ankle 2.28 (1.00, 5.19)	0.01	0.21	0.14
tender joints*	MTP1 2.82 (1.37, 5.81)	0.01	0.42	0.30
Serum urate level > 6mg/dl (0.36mmol/l)	3.35 (1.57, 7.15)	0.002	1.21	1.43
US double contour sign	7.23 (3.47, 15.04)	<0.0001	1.98	1.40
Xray erosion or cyst	2.49 (1.26, 4.90)	0.009	0.91	0.65

* Ref. category is joint proximal to ankle

Logistic regression model details: LR Chi-square 305 (df 11), p<0.001; c-statistic 0.93
The regression coefficients show the strength of the association between the variable and having gout (values > 0 indicate a positive association) in the original units of the variable. The shrinkage factor is applied to prevent over-estimating the predictive ability of the model (see text for more details).

OR = odds ratio; MTP1 = first metatarsophalangeal joint; US = ultrasound

Figure 1. Association of highest serum urate (mg/dl) with gout. The size of the bars represent the ratio of the proportion of gout/non-gout with each level of serum urate. The bar labels are the actual observed proportion in the gout or non-gout groups with each level of serum urate.



John Wiley & Sons, Inc.

References

1.

4.

9.

Neogi T. Clinical practice. Gout. N Engl J Med. 2011;364(5):443-52.

- 2. Trifiro G, Morabito P, Cavagna L, Ferrajolo C, Pecchioli S, Simonetti M, et al. Epidemiology of gout and hyperuricaemia in Italy during the years 2005-2009: a nationwide population-based study. Ann Rheum Dis. 2013;72(5):694-700.
 - Wallace SL, Robinson H, Masi AT, Decker JL, McCarty DJ, Yu TF. Preliminary criteria for the classification of the acute arthritis of primary gout. Arthritis Rheum. 1977;20(3):895-900.
 - Dalbeth N, Fransen J, Jansen TL, Neogi T, Schumacher HR, Taylor WJ. New
 classification criteria for gout: a framework for progress. Rheumatology (Oxford).
 2013;52(10):1748-53.
- 5. Malik A, Schumacher HR, Dinnella JE, Clayburne GM. Clinical diagnostic criteria for gout: comparison with the gold standard of synovial fluid crystal analysis. JCR: Journal of Clinical Rheumatology. 2009;15(1):22-4.
- 6. Janssens HJEM, Janssen M, van de Lisdonk EH, Fransen J, van Riel PLCM, van Weel C. Limited validity of the American College of Rheumatology criteria for classifying patients with gout in primary care. Ann Rheum Dis. 2010;69(6):1097-102.
 7. Johnson SR, Goek ON, Singh_Grewal D, Vlad SC, Feldman BM, Felson DT, et al. Classification criteria in rheumatic diseases: a review of methodologic properties. Arthritis Rheum. 2007;57(7):1119-33.
 - 8. Taylor WJ, Robinson PC. Classification Criteria: Peripheral Spondyloarthropathy and Psoriatic Arthritis. Curr Rheumatol Rep. 2013;15(317):1-7.
 - Rudwaleit M, Taylor WJ. Classification criteria for psoriatic arthritis and ankylosing spondylitis/axial spondyloarthritis. Bailliere's Best Practice and Research Clinical Rheumatology. 2010;24(5):589-604.

- Bossuyt PM, Reitsma JB, Bruns DE, Gatsonis CA, Glasziou PP, Irwig LM, et al. Towards complete and accurate reporting of studies of diagnostic accuracy: The STARD initiative. Clin Chem. 2003;49(1):1-6.
- 11. Berendsen D, Jansen TL, Taylor W, Neogi T, Fransen J, Pascual E, et al. A critical appraisal of the competence of crystal identification by rheumatologists (abstract). Ann Rheum Dis. 2013;72(Suppl 3):981.
- 12. Prowse RL, Dalbeth N, Kavanaugh A, Adebajo AO, Gaffo AL, Terkeltaub R, et al. A delphi exercise to identify characteristic features of gout opinions from patients and physicians, the first stage in developing new classification criteria. J Rheumatol. 2013;40(4):498-505.
- 13. Steyerberg EW, Eijkemans MJ, Harrell FE, Jr., Habbema JD. Prognostic modelling with logistic regression analysis: a comparison of selection and estimation methods in small data sets. Stat Med. 2000;19(8):1059-79.
- Steyerberg EW, Eijkemans MJC, Habbema JDF. Application of Shrinkage
 Techniques in Logistic Regression Analysis: A Case Study. Statistica Neerlandica.
 2001;55(1):76-88.
 - 15. Rubin DB. Multiple Imputation for Nonresponse in Surveys. New York: John Wileyand Sons; 2004.
 - 16. Lanza ST, Collins LM, Lemmon DR, Schafer JL. PROC LCA: A SAS Procedure for Latent Class Analysis. Struct Equ Modeling. 2007;14(4):671-94.
- 17. Collins LM, Lanza ST. Latent Class and Latent Transition Analysis. Hoboken, NJ: John Wiley & Sons, Inc.; 2010.
- 18. Lanza ST, Bray BC. Transitions in Drug Use among High-Risk Women: An
 Application of Latent Class and Latent Transition Analysis. Adv Appl Stat Sci. 2010;3(2):203-35.

John Wiley & Sons, Inc.

Hall AP, Barry PE, Dawber TR, McNamara PM. Epidemiology of gout and hyperuricemia. Am J Med. 1967;42:27-37.

19.

Acce

20. Roddy E. Revisiting the pathogenesis of podagra: why does gout target the foot? J Foot Ankle Res. 2011;4(1):13.

John Wiley & Sons, Inc.



258x169mm (150 x 150 DPI)

Accepte