- 1 Low-cost sensor networks and land-use regression: interpolating nitrogen dioxide
- 2 concentration at high temporal and spatial resolution in Southern California
- 3 Lena Weissert<sup>1,3</sup> a, Kyle Alberti<sup>2</sup>, Elaine Miles<sup>2</sup>, Georgia Miskell<sup>1b</sup>, , Brandon Feenstra<sup>4</sup>, Geoff
- 4 S Henshaw<sup>2</sup>, Vasileios Papapostolou<sup>4</sup>, Hamesh Patel<sup>2</sup>, Andrea Polidori<sup>4</sup>, Jennifer A Salmond<sup>3</sup>,
- 5 David E Williams<sup>1,\*</sup>
- 6 \*Email <u>david.williams@auckland.ac.nz</u> ph +64 9 923 987**7**
- 7 1. School of Chemical Sciences and MacDiarmid Institute for Advanced Materials and
- 8 Nanotechnology, University of Auckland, Private Bag 92019, Auckland 1142, New Zealand
- 9 2. Aeroqual Ltd, 460 Rosebank Road, Avondale, Auckland 1026, New Zealand
- 3. School of Environment, University of Auckland, Private Bag 92019, Auckland 1142, New
- 11 Zealand
- 4. South Coast Air Quality Management District, 21865 Copley Drive, Diamond Bar, CA
- 13 91765, USA

15

#### Abstract

- 16 The development of low-cost sensors and novel calibration algorithms offer new opportunities
- to supplement existing regulatory networks to measure air pollutants at a high spatial resolution
- and at hourly and sub-hourly timescales. We use a random forest model on data from a network
- 19 of low-cost sensors to describe the effect of land use features on local-scale air quality, extend
- 20 this model to describe the hourly-scale variation of air quality at high spatial resolution, and
- show that deviations from the model can be used to identify particular conditions and locations

<sup>&</sup>lt;sup>a</sup> Present address: Aeroqual Ltd, 460 Rosebank Road, Avondale, Auckland 1026, New Zealand

<sup>&</sup>lt;sup>b</sup> Present address: Trustpower, 108 Durham St, Tauranga, New Zealand

- where air quality differs from the expected land-use effect. The conditions and locations under
- 23 which deviations were detected conform to expectations based on general experience.

Keywords

24

25

27

26 Air quality sensor network; land-use regression; nitrogen dioxide; ozone

#### Introduction

28

29

30

31

32

33

34

35

36

37

38

39

40

41

42

43

44

45

46

47

48

49

50

51

52

The South Coast Air Basin is one of the most polluted air basins in the United States (Epstein et al., 2017). The pollution problem in this region is driven by high emissions, unfavourable meteorological conditions (low wind speed, strong temperature inversions, abundant sunshine, infrequent rainfall), sea breezes and complex terrain that limits pollutant dispersion (South Coast AQMD, 2016). Spatially and temporally dense information about local scale air pollution is necessary to mitigate air pollution effectively (Vizcaino and Lavalle, 2018). While regulatory air quality monitoring networks offer important insights about long-term air quality trends, the data must be supplemented with additional measurements and models to obtain geographically more detailed air pollution information (Li et al., 2019a). This is of importance given that air pollutant concentrations can vary considerably over small distances (Kumar et al., 2015; Weissert et al., 2019a). Association of average pollutant concentration with land use variables (e.g. distance to major roads, length of major roads within different buffers, bus stops) is a frequently used approach to model time-averaged pollutant concentrations with high spatial resolution (Hoek et al., 2008). A limitation of land use regression (LUR) models is the risk of overfitting the data when only few measurement sites are used to train the model. Further, LUR modelling is based on the assumption that relationships between air pollution and predictor variables are linear and that there are no interaction effects between different predictors. Other algorithms that remove some of these limitations (e.g. Generalized Additive Model (GAM), Least Absolute Shrinkage and Selection Operator (LASSO)) have been used to fit a land use model to pollutant concentrations (Chen et al., 2019). Some studies have also used machine learning algorithms such as Random Forest (RF) (Brokamp et al., 2017; Hu et al., 2017; Zhan et al., 2018). LUR models are usually developed from dense diffusion tube monitoring over a few weeks during different seasons and lack temporal resolution at the hourly or sub-hourly scale. To overcome

models (Masiol et al., 2018; Miskell et al., 2018b; Son et al., 2018; Yeganeh et al., 2018). The development of low-cost sensors has created new opportunities for air quality measurements and modelling. If deployed in dense networks, low-cost sensors have the potential to provide near real-time measurements of pollutants at a spatial resolution representative of the neighbourhood scale. They can offer insights into the influence of local pollution sources at different temporal and spatial scales that may not be detected by the usually sparsely distributed regulatory monitoring networks (Feinberg et al., 2019; Li et al., 2019b; Popoola et al., 2018; Weissert et al., 2019a). Hence, the increasingly available data from lowcost sensor networks has led to new research aimed at combining continuous measurements obtained from a low-cost sensor network with land use data to get spatially and temporally dense air pollution information (Deville Cavellin et al., 2016; Lim et al., 2019; Masiol et al., 2019; Miskell et al., 2018b; Schneider et al., 2017). In Montreal and Vancouver, Canada, mobile measurements with low-cost sensors were used to map air pollutants during different seasons (Deville Cavellin et al., 2016) and times of the day (Miskell et al., 2018b).. Schneider et al. (2017) combined air quality data obtained from a low-cost sensor network (24 units) with an urban-scale air quality dispersion model to map nitrogen dioxide (NO<sub>2</sub>) concentrations at near real-time. A network of ten low-cost sensors was used in New York to develop 24 LUR models representative of each hour of the day (Masiol et al., 2019). In a recently published pilot study, we presented another approach to combine NO<sub>2</sub> concentrations obtained from a microscale low-cost sensor network (eight sensors along a 2 km length of a heavily-trafficked road with mixed commercial and residential land use) with land use information to identify site and time specific effects of urban design features that disproportionately contribute to population exposure (Weissert et al., 2019a).

this, LUR models have been combined with temporally variable predictors to obtain hourly

53

54

55

56

57

58

59

60

61

62

63

64

65

66

67

68

69

70

71

72

73

74

75

76

Such attempts to fuse land-use and sensor network data face the challenge of demonstrating plausibility of data from low-cost sensor networks (Williams, 2019). A considerable amount of research has focused on sensor performance. A specific issue is drift of sensor signals over time (Clements et al., 2017). Thus, an increasing number of researchers are focusing on developing procedures that allow remote sensor calibrations (Delaine et al., 2019), which is critical for the long-term deployment of large low-cost sensor networks. In our recent work, we developed calibration and remote drift detection procedures for ozone (O<sub>3</sub>) and NO<sub>2</sub> sensors deployed in hierarchical networks consisting of a few well-maintained regulatory sites and a large number of low-cost sensors, distributed across a region. The approaches were tested with success for networks installed in the Lower Fraser Valley, Canada, and in Southern California (Miskell et al., 2019; Miskell et al., 2016; Miskell et al., 2018a; Weissert et al., 2019a) . Correction of sensor drift over time, and of offset errors, has been comprehensively demonstrated. The corrected sensor data for Southern California provide a spatially and temporally dense data set of O<sub>3</sub> and NO<sub>2</sub> concentrations that can claim reliability with a root mean-square error (RMSE) of 5.4 and 7.4 ppb, respectively, over the several months of the study. Here, first we use this dataset to develop a land-use model for concentrations averaged over two months, and then apply the simple approach described by Weissert et al. (2019a) to model concentrations on an hourly time-scale, both for O<sub>3</sub> and NO<sub>2</sub>. Then, we use an analysis of differences between measured and modelled concentrations to identify local urban conditions that are poorly captured by the static land use model. We show that these deviations have reasonable explanations which in turn reinforces confidence in the original dataset (Williams, 2019).

99

100

77

78

79

80

81

82

83

84

85

86

87

88

89

90

91

92

93

94

95

96

97

98

### Methods

## Sensor network

We used the micro air quality monitoring instruments (model AQY) from Aeroqual Ltd., Auckland, New Zealand, which are described in detail in Weissert et al. (2019a). The data correction procedures for electrochemical NO<sub>2</sub> sensors to account for interference by ozone and for offset errors have been comprehensively described elsewhere (Miskell et al., 2019; Miskell et al., 2018a, Weissert et al., 2019b, c). The cross-sensitivity to NO of electrochemical NO<sub>2</sub> sensors is small and can be ignored (Mead et al., 2013; Popoola et al., 2018). The low-cost sensor network was deployed in the Inland Empire in Southern California (Fig. 1).

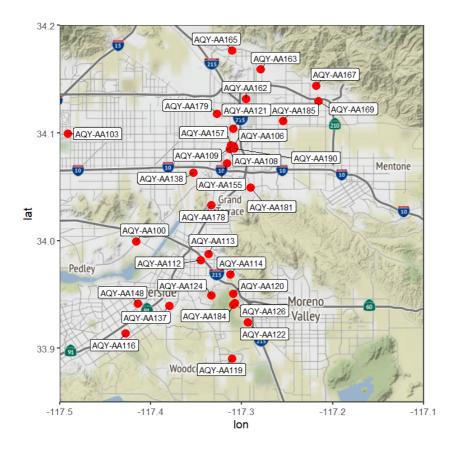


Figure 1. Low-cost sensor sites used for this study (n = 31).

For the model building, we used data from 31 low-cost sensor sites during April and May 2018, when most data were available. Pollutant concentrations (O<sub>3</sub> and NO<sub>2</sub>) were averaged across the two months to develop the 'average' model.

## Predictor variables

Publicly accessible land use data from Open Street Map and traffic data (Caltrans, 2019) were used as predictor variables. Altitude data was extracted from the Shuttle Radar Topography Mission (SRTM) 90m Digital Elevation data (http://srtm.csi.cgiar.org/). We used three variable selection methods to assess the effect of different predictor numbers offered to the model. First, we used all available predictors (Chen et al., 2019), second we optimized buffer distances (Su et al., 2009; Vizcaino and Lavalle, 2018) and third, we removed variables that did not follow the expected direction of effect (Beelen et al., 2013; Vizcaino and Lavalle, 2018). To optimize buffer distances, the correlations between the pollutant concentrations and the predictor variable at each buffer distance were calculated and the one with the highest value of correlation was used in the model.

Table 1. Predictor variables used in the model. Buffer size refers to the radius around the studysite.

Predictor variables	Variable code	Unit	Buffer size (m)
Altitude	Elevation	m	
Coordinates of the low-cost			
instrument site	Lat/Long	-	
Length of all main roads			24, 50, 100, 200,
within the buffer circle	MAJORROADLENGTH	m	300, 500, 1000
Length of all roads within			24, 50, 100, 200,
the buffer circle	ROADLENGTH	m	300, 500, 1000
Inverse distance to the			
nearest main road	DISTINVNEAR1	$m^{-1}$	
Truck traffic	TRUCK_AADT	veh day <sup>-1</sup>	
Vehicle traffic	VEH_AADT	veh day <sup>-1</sup>	

# Model building and validation

We used a random forest (RF) model. RF models have successfully been used in previous studies aiming to predict NO<sub>2</sub> concentrations using land use (Araki et al., 2018; Chen et al., 2019; Hu et al., 2017; Zhan et al., 2018). This approach was chosen to minimise the risk of overfitting given that there are relatively few monitoring sites and also to capture non-linear relationships observed between air pollutant concentrations and predictor variables (Araki et al., 2018; Chen et al., 2019; Vizcaino and Lavalle, 2018). RF models are bagged decision tree models, where each tree consists of a random subset of predictor variables from the training dataset and where the final output is the average of multiple decision trees (Breiman, 2001; Grange et al., 2018; Vizcaino and Lavalle, 2018).

We used the caret package in R (v3.5.3) to develop the RF (Kuhn, 2019). Ideally, the data would be split into a training (80%) set, which is used to develop the model, and a test (20%) set, which is used to evaluate the performance of the model (hold-out validation). However, given the small sample size of sites, we decided to use all sites to develop the model. Thus, we could not verify the performance of the model on held-out test data. Therefore, the model developed for the monitoring sites may not be representative of other sites. However, the focus of this paper is to assess local effects that result in deviations from the average modelled concentrations, which would not be as affected by the lack of test data. The model is evaluated using a 10-fold cross-validation for resampling where the root mean square error (RMSE) is taken as the metric to measure model performance.

# Fusion of the RF model with hourly-averaged data

To build the model for the temporal variation at the hourly-averaged time-scale, we used the approach described in Weissert et al. (2019a). In brief, we assume that the modelled

concentrations  $(\bar{C}_{RF,k})$  are linearly related to the hourly-averaged low-cost instrument data  $(y_k)$ 

for any given hour on any given day (l).

155 
$$y_{k,l} = \hat{a}_{1,l} \, \bar{C}_{RF,k} + e_{l,k}$$
 (1)

where  $\hat{a}_{1,l}$  is derived from a least-square regression of eq. 1. An analysis of  $e_l$  at the different

low-cost sensor sites, k, was then used to assess local effects that are not captured by the RF

model.

159

160

161

162

163

164

165

157

158

# Results and Discussion

Measured pollutant concentrations

Concentrations are expressed as mixing ratios: parts-per-billion (10<sup>9</sup>) by volume, ppb. Figure

2 shows boxplots for the O<sub>3</sub> and NO<sub>2</sub> concentrations measured at the different low-cost sensor

sites across the study period, showing that the intra-site variability tends to be larger than the

variability between sites for both pollutants. Maximum 8-hour O<sub>3</sub> was 120 ppb (site 184). The

highest 1-hour average NO<sub>2</sub> concentrations were recorded at site 124 (116 ppb).

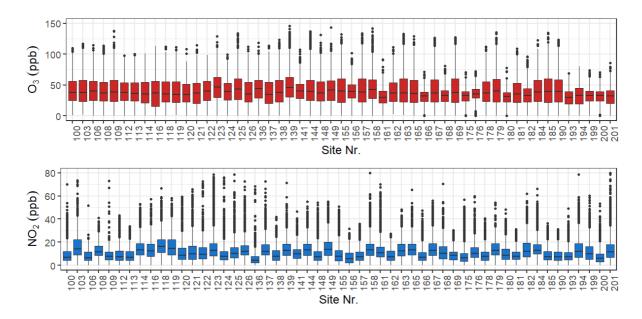


Figure 2. Boxplot for NO<sub>2</sub> and O<sub>3</sub> concentrations measured at the low-cost sensor sites (x-axis) from April to May 2018. The line denotes the median value. The upper and lower hinges represent the 25th and 75th percentiles. The whiskers extend from the hinge 1.5 times the interquartile range. Outliers are shown as dots. The location of the sites and site numbers are shown in figure 1.

Figure 3 shows the spatial variability of mean O<sub>3</sub> and NO<sub>2</sub> concentrations. Mean O<sub>3</sub> concentrations were not highly variable between locations across the region. However, the range of concentrations experienced between locations did differ significantly (figure 2). The mean NO<sub>2</sub> concentration was highly spatially variable across the region. The results suggest higher O<sub>3</sub> and NO<sub>2</sub> concentrations north of Riverside along the mountain range. At other sites, the two pollutants show the expected opposite pattern with higher NO<sub>2</sub> concentrations and lower O<sub>3</sub> concentration in the south west (SW) direction of Riverside. Although at individual sites the NO<sub>2</sub> concentration showed an irregular temporal variation (Weissert et al. 2019b), on average, with more variability for NO<sub>2</sub>, the two pollutants showed a simple diurnal variation

with the lowest value close to zero: figure 4. For a regular diurnal variation with minimum zero, eq 1 would apply exactly.

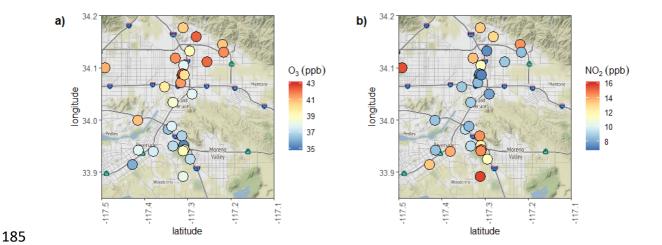


Figure 3. Average measured  $O_3$  (left) and  $NO_2$  (right) concentrations at the low-cost sensor sites.

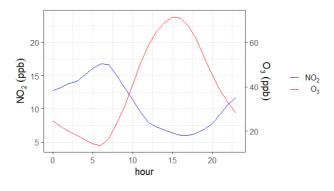


Figure 4. Mean diurnal variation of O<sub>3</sub> (red) and NO<sub>2</sub> (blue), averaged over all sites and days.

# Model results

Table 2 shows a summary of the resampling results for different predictor selection approaches (1: all predictors, 2: predictors with optimized buffers, 3: predictors with optimized buffers and

that follow the expected direction of effect). It shows that the model performed well on the data, with a slightly better performance when using all predictors. When applied to all sensor sites, the sample Pearson correlation coefficient,  $R^2$ , between the modelled and measured  $O_3$  and  $NO_2$  concentrations was 0.73 (RMSE = 1.8 ppb) and 0.93 (RMSE = 1.3 ppb), respectively.

Table 2. Summary of the model performance for different predictor selections (1: all predictors, 2: predictors with optimized buffers, 3: predictors with optimized buffers and that follow the expected direction of effect).  $m_{try}$  is the number of variables randomly sampled as candidates.

Approach	Training data $(n = 31)$					
	O <sub>3</sub>			NO <sub>2</sub>		
	$R^2$	RMSE	$m_{\mathrm{try}}$	$R^2$	RMSE	$m_{\mathrm{try}}$
1	0.70	1.14	11	0.71	1.82	20
2	0.71	1.14	5	0.66	1.95	2
3	NA	NA	NA	0.66	1.95	2

Figure 5 shows the variable importance derived from the RF suggesting that location (latitude) is the most important predictor for O<sub>3</sub>, followed by average truck traffic and the inverse distance from the nearest main road. The spatial variability of the measured pollutant concentrations confirms the higher O<sub>3</sub> concentrations at higher latitudes at the bottom of the mountain range (Fig. 3). NO<sub>2</sub> concentrations were largely dependent on the main road length within 1 km, inverse distance to main road and elevation with a tendency for higher concentrations measured at higher altitudes.

Whilst the inverse distance to the nearest main road was important, there was a lot of scatter and the relationship with NO<sub>2</sub> concentrations was weak ( $R^2 < 0.1$ ) and followed the opposite direction of effect. While predictors following an unexpected direction of effect are excluded in standard linear regression models, RF models may also include predictors with counter-

intuitive effects, for example, to compensate for over or under predictions by other predictors (Chen et al., 2019).

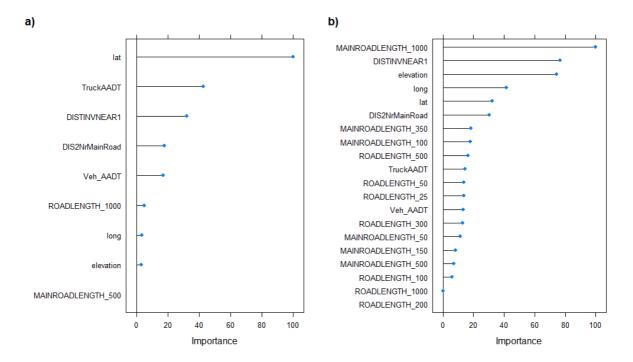


Figure 5. Scaled variable importance (%) plot for the final RF model for a) O<sub>3</sub>, b) NO<sub>2</sub>. The variables are listed in order of importance from top to bottom.

# Temporal variation results

Figure 6 shows the hourly-averaged O<sub>3</sub> concentrations measured at the low-cost sensor sites against the modelled and temporally updated O<sub>3</sub> concentrations. Figure 6 shows that across the entire study period, the model captured well the temporal variation measured at the low-cost sensor sites. This is to be expected since at most sites and times the variation was a regular diurnal cycle with the lowest value close to zero; hence, the hourly variation would be simply related to the mean.

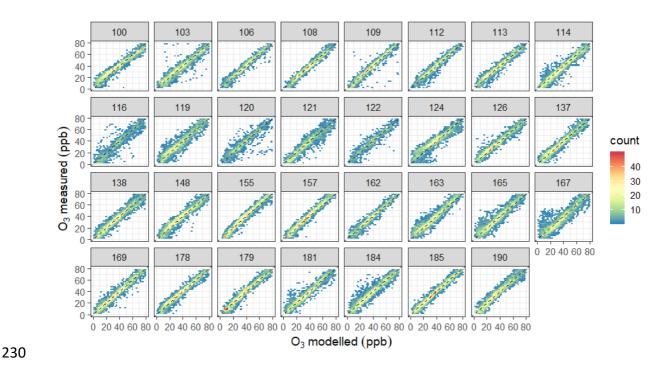


Figure 6. Hexbin plots of hourly averaged  $O_3$  concentrations at the low-cost sensor sites against the modelled  $O_3$  concentrations (equation 1). The dashed line is the 1:1 line.

Figure 7 shows the same figure for NO<sub>2</sub>, with the measured NO<sub>2</sub> concentrations on the y-axis and the modelled NO<sub>2</sub> concentrations on the x-axis. The model captured the overall temporal variability well at most sites, however some deviations can be observed. At site 121, for example, the model was not able to capture the NO<sub>2</sub> concentrations measured at this site. Likewise, some high concentrations were missed at site 124.

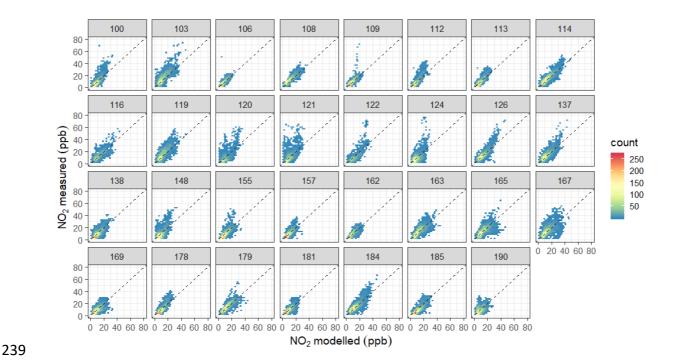


Figure 7. Hexbin plots of hourly averaged NO<sub>2</sub> concentrations at the low-cost instrument sites against the modelled NO<sub>2</sub> concentrations (equation 1). The dashed line is the 1:1 line.

# Analysis of local effects

Figure 8 shows the unexplained variance for each site as a fraction of the total unexplained variance.  $O_3$  concentrations were generally well captured, which is partly due to the lower spatial variability of  $O_3$ . Sites where the model did not predict temporal  $O_3$  concentrations as well include sites 120, 124,116 and 167. The unexplained variance was slightly higher for  $NO_2$  (Fig. 8b) and higher for sites 121, 120, 124 and 167. The spatial variation of the unexplained variance is shown in figure 8c - e. As indicated by the larger unexplained variance (expressed as the ratio of the mean sum of squared error, sse, at the particular site divided by the mean sum of squared error at all sites, sse site/sse total) the model did not perform as well for  $O_3$  and  $NO_2$  for sites close to the mountain range, north and south of the valley. In an effort to analyse and discuss local effects that may have contributed to unpredicted variations, we also plotted

the mean difference term (measured – modelled concentrations) across different wind direction-speed bins (Fig. 9). Some examples for local effects are discussed below.

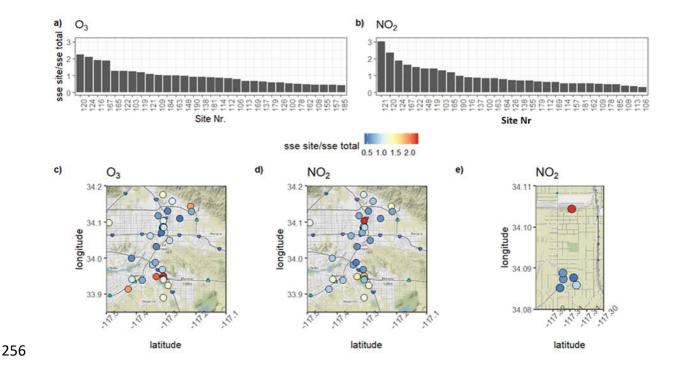


Figure 8. a) - b)  $O_3$  and  $NO_2$  variation not explained by the model at the different low-cost instrument sites; mean sum of squared error (sse) at the particular site divided by mean sum of squared error at all sites (sse total). c) - e) maps at different scales to illustrate the spatial variability of the unexplained  $O_3$  and  $NO_2$  variation. The location of the sites and site numbers are shown in figure 1.

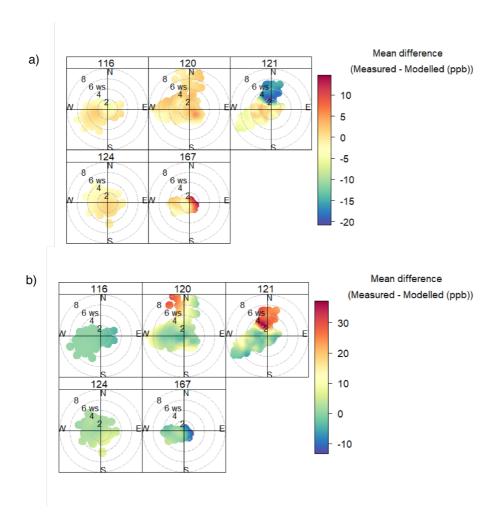


Figure 9. Polar plots showing the mean difference between the measured and modelled  $O_3$  (a) and  $NO_2$  (b) concentrations divided into different wind direction and wind speed (ws / m s<sup>-1</sup>) bins.

Site 116 is located SW of Riverside at a high school. The O<sub>3</sub> model tended to slightly overestimate O<sub>3</sub> concentrations at this site, particularly when wind speed was low. This suggests that O<sub>3</sub> concentrations at this school are lower than expected for sites at this latitude and sites with similar traffic and road patterns. The reason may be hyper-local traffic patterns around the site that may be leading to local O<sub>3</sub> titration with vehicle-emitted nitric oxide (NO). Site 120 is located in a residential area, east of the Sycamore Canyon Wilderness Park. Thus, the main road length within 1 km, one of the most important predictors for NO<sub>2</sub>, was relatively

small. However, the site was also S/SW of a multi-lane motorway and it is likely that high NO<sub>2</sub> and low O<sub>3</sub> concentrations were measured when the site was downwind from the motorway, as can be seen in figure 9. Another site that showed distinct differences between measured and modelled NO<sub>2</sub> concentrations was site 121. The polar-plot in figure 9 reveals that this was particularly the case when there was a north wind direction. This site was located just south of the San Bernardino Santa Fe depot, a major road and rail transport hub (Fig. 8e), which resulted in high NO<sub>2</sub> concentrations at this site. Site 124 was located 100 m west from a major road, but measured O<sub>3</sub> concentrations were higher and NO<sub>2</sub> lower than typically expected within close proximity to a main road. The site is located within a golf-course, suggesting perhaps that nitric oxide (NO) was scavenged by grass and vegetation so O<sub>3</sub> titration from vehicle traffic may have been lower. Site 167 is 400 m south-west from the San Bernardino National Forest, which explains the higher than modelled O<sub>3</sub> concentrations and lower than modelled NO<sub>2</sub> concentrations associated with wind from the forested area being lower in NO. To further assess the temporal differences between modelled and measured concentrations, we examined the time-series for April for the sites with disproportionately high unexplained variances (see Figure 10). For O<sub>3</sub>, the temporally updated model followed the measured O<sub>3</sub> concentrations relatively well, although some deviations were observed between the 9th and 14<sup>th</sup> of April at site 121 where O<sub>3</sub> concentrations were lower than modelled. This site also showed large deviations for NO<sub>2</sub> for the same time period, suggesting the local emission sources may have changed during this time period. Wind data for this period showed a change in wind direction from dominating south-westerlies to northerlies. Thus, the influence from the

275

276

277

278

279

280

281

282

283

284

285

286

287

288

289

290

291

292

293

294

295

296

297

298

299

higher than modelled. Similarly, site 167 showed larger deviations between the  $9^{th}$  and  $14^{th}$  of

April, when measured air was mostly coming from the San Bernardino National Forest.

train depot north of this site would therefore have been stronger between the 9th and 14th of

April, particularly on the 9<sup>th</sup> and 13<sup>th</sup> when measured NO<sub>2</sub> concentrations were considerably

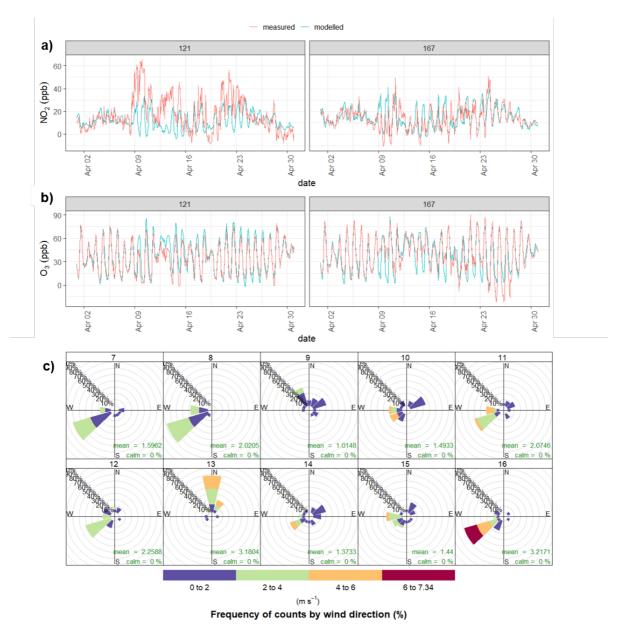


Figure 10. a) hourly measured (red) and modelled (blue)  $NO_2$  concentrations and b) hourly measured (red) and modelled (blue)  $O_3$  concentrations during April, c) frequency of wind speed and wind direction in different wind speed and wind direction categories at site 121 between the  $7^{th}$  and  $16^{th}$  of April (the panels are different days).

Temporally variable pollution maps

Finally, the presented approach allows mapping pollutant concentrations for any given day and hour measurement data are available. An example for predicted NO<sub>2</sub> concentrations for 10/05/2018 at 07:00 local time is shown in figure 11. Given that the prediction success of the model developed here is limited due to the relatively small number of sites, the figure is focused on an area where measurements were relatively dense and the predictions likely more representative. It shows the expected high NO<sub>2</sub> concentrations during the morning rush hour. Spatially, NO<sub>2</sub> concentrations were higher north and south of the study area, for which higher altitude was the main driver in the model. The likely explanation is the reaction with vehicleemitted NO of ozone transported down the valley sides from higher in the troposphere: the ozone concentration variation shown in figure 8 is consistent with this. Although the local spatial variations seem small, they are of importance in relation to the emerging understanding of the consequences for health of increases in  $NO_2$  concentration on the scale of 5-10 ppb and in relation to suggestions of the need for quantification of long-term health effects of annual mean concentrations > 10 ppb (WHO, 2013). Estimates of potential harm are very sensitive to the thresholds chosen (European Environment Agency, 2019). In the analysis of epidemiological effects, having data at sufficiently fine spatial and temporal scale is important (Wei et al., 2019). The spatial coverage and prediction power of the model may be improved by supplementing low-cost sensor networks with diffusion tube campaigns involving local communities or schools. The combination of the two approaches would not only show the general land-use effects on average air quality but also would show the temporal variation as well as the particular, time-dependent effect of local urban design features not captured by the land use model. This would then allow mapping pollutant concentrations with more confidence for any given day and hour at a neighbourhood scale and offer insights about pollution hotspots and their temporal variation.

308

309

310

311

312

313

314

315

316

317

318

319

320

321

322

323

324

325

326

327

328

329

330

331

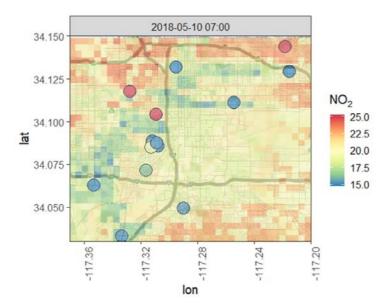


Figure 11. Predicted NO<sub>2</sub> concentrations (ppb) at 500 m resolution for northern Riverside, where the low-cost sensor network was the densest network in the region, with measured NO<sub>2</sub> concentrations at the sensor sites superimposed: 10<sup>th</sup> May 2018 at 07:00. The dark lines are major roads.

## **Conclusions**

This paper has presented results of a RF model to predict concentration values for  $O_3$  and  $NO_2$  based on LUR and a low-cost sensor network that was deployed in the Inland Empire region in Southern California. A previously described procedure was used to remotely calibrate the low-cost sensors using data from the more sparsely distributed regulatory network. We combined land use information and an RF model with hourly low-cost sensor data to identify local effects on  $O_3$  and  $NO_2$  concentrations at a high temporal resolution. The mean RF models performed well for  $O_3$  and  $NO_2$  concentrations ( $R^2 = 0.73/RMSE = 1.8$  ppb and  $R^2 = 0.93/RMSE = 1.3$  ppb, respectively) at the low-cost sensor sites. The model consistency over the several months of the study gave support to the effectiveness of the sensor drift and offset correction methods. The mean modelled pollutant concentrations were successfully updated

hourly using the low-cost instrument data. The model for O<sub>3</sub> combined with the low-cost instrument data captured the spatial and temporal variation well. For NO<sub>2</sub>, deviations from the model highlighted particular urban features, not accounted for by the general land-use modelling, that under particular circumstances resulted in significantly increased pollutant concentration that would be cause for concern. These locations and circumstances could be related back to specific and understandable local effects. The combination of sensor data at high time resolution and an average land-use model proved to be an effective and simple way to highlight the spatial and temporal distribution of local effects that may disproportionately contribute to pollutant concentrations. The findings may be associated with different meteorological conditions (e.g. higher pollutant concentrations expected for particular wind directions) offering support for local pollution alerts. If supplemented with more dense measurements, for example using diffusion tube campaigns, the model would allow for mapping pollutant concentrations for any given day and hour, which may be updated in near real-time as long as measurement data were available.

## Acknowledgement

This work was funded by the New Zealand Ministry for Business, Innovation and Employment, contract UOAX1413. This work was performed in collaboration with the Air Quality Sensor Performance Evaluation Center (AQ-SPEC) at the South Coast Air Quality Management District (South Coast AQMD). The authors would like to acknowledge the work of Mr. Berj Der Boghossian for his technical assistance with deploying AQY sensor nodes. The authors would like to acknowledge the work of the South Coast AQMD Atmospheric Measurements group of dedicated instrument specialists that operate, maintain, calibrate, and repair air monitoring instrumentation to produce regulatory-grade air monitoring data. DEW acknowledges the support of a fellowship program at the Institute of Advanced Studies, Durham University, UK.

6. Competing interest
-----------------------

376 LW, EM, KA and GSH are employees of Aeroqual Ltd, manufacturer of the sensor nodes used in the study. GSH and DEW are founders and shareholders in Aeroqual Ltd. 377

378

379

394

395

396

397

## References

Araki, S., Shima, M., Yamamoto, K., 2018. Spatiotemporal land use random forest model for 380 estimating metropolitan NO<sub>2</sub> exposure in Japan. Sci Total Environ 634, 1269-1277. 381 Beelen, R., Hoek, G., Vienneau, D., Eeftens, M., Dimakopoulou, K., Pedeli, X., Tsai, M.-Y., 382 383 Künzli, N., Schikowski, T., Marcon, A., Eriksen, K.T., Raaschou-Nielsen, O., Stephanou, E., Patelarou, E., Lanki, T., Yli-Tuomi, T., Declercq, C., Falq, G., 384 Stempfelet, M., Birk, M., Cyrys, J., von Klot, S., Nádor, G., Varró, M.J., Dėdelė, A., 385 386 Gražulevičienė, R., Mölter, A., Lindley, S., Madsen, C., Cesaroni, G., Ranzi, A., 387 Badaloni, C., Hoffmann, B., Nonnemacher, M., Krämer, U., Kuhlbusch, T., Cirach, M., 388 de Nazelle, A., Nieuwenhuijsen, M., Bellander, T., Korek, M., Olsson, D., Strömgren, 389 M., Dons, E., Jerrett, M., Fischer, P., Wang, M., Brunekreef, B., de Hoogh, K., 2013. Development of NO<sub>2</sub> and NO<sub>x</sub> land use regression models for estimating air pollution 390 exposure in 36 study areas in Europe - The ESCAPE project. Atmospheric 391 392 Environment 72, 10-23. 393

Breiman, L., 2001. Random Forests. Machine Learning 45, 5 - 32.

Brokamp, C., Jandarov, R., Rao, M.B., LeMasters, G., Ryan, P., 2017. Exposure assessment models for elemental components of particulate matter in an urban environment: A comparison of regression and random forest approaches. Atmos Environ (1994) 151, 1-11.

- 398 Caltrans (2019). Caltrans GIS Data Truck Volumes AADT. https://gisdata-
- caltrans.opendata.arcgis.com/datasets/dfe7fd95282946db98145e9bcaf710fb\_0,
- 400 Accessed: September, 2019).
- 401 Chen, J., de Hoogh, K., Gulliver, J., Hoffmann, B., Hertel, O., Ketzel, M., Bauwelinck, M.,
- van Donkelaar, A., Hvidtfeldt, U.A., Katsouyanni, K., Janssen, N.A.H., Martin, R.V.,
- Samoli, E., Schwartz, P.E., Stafoggia, M., Bellander, T., Strak, M., Wolf, K., Vienneau,
- D., Vermeulen, R., Brunekreef, B., Hoek, G., 2019. A comparison of linear regression,
- regularization, and machine learning algorithms to develop Europe-wide spatial models
- of fine particles and nitrogen dioxide. Environ Int 130, 104934.
- Clements, A.L., Griswold, W.G., Rs, A., Johnston, J.E., Herting, M.M., Thorson, J., Collier-
- Oxandale, A., Hannigan, M., 2017. Low-Cost Air Quality Monitoring Tools: From
- 409 Research to Practice (A Workshop Summary). Sensors (Basel) 17.
- 410 Delaine, F., Lebental, B., Rivano, H., 2019. In Situ Calibration Algorithms for Environmental
- Sensor Networks: A Review. IEEE Sensors Journal 19, 5968-5978.
- Deville Cavellin, L., Weichenthal, S., Tack, R., Ragettli, M.S., Smargiassi, A., Hatzopoulou,
- 413 M., 2016. Investigating the Use Of Portable Air Pollution Sensors to Capture the
- Spatial Variability Of Traffic-Related Air Pollution. Environ Sci Technol 50, 313-320.
- 415 Epstein, S.A., Lee, S., Katzenstein, A.S., Carreras-Sospedra, M., Zhang, X., Farina, S.C.,
- Vahmani, P., Fine, P.M., Ban-Weiss, G., 2017. Air-quality implications of widespread
- adoption of cool roofs on ozone and particulate matter in southern California. PNAS
- 418 114, 8991-8996.
- 419 European Environment Agency, 2019. Air Quality in Europe 2019 Report. Publications
- 420 Office of the European Union, Luxembourg.
- 421 Feinberg, S.N., Williams, R., Hagler, G., Low, J., Smith, L., Brown, R., Garver, D., Davis,

422 M., Morton, M., Schaefer, J., Campbell, J., 2019. Examining spatiotemporal 423 variability of urban particulate matter and application of high-time resolution data from a network of low-cost air pollution sensors. Atmospheric Environment 213, 579-424 425 584. 426 Grange, S.K., Carslaw, D.C., Lewis, A.C., Boleti, E., Hueglin, C., 2018. Random forest 427 meteorological normalisation models for Swiss PM<sub>10</sub>; trend analysis. Atmospheric 428 Chemistry and Physics 18, 6223-6239. 429 Hoek, G., Beelen, R., de Hoogh, K., Vienneau, D., Gulliver, J., Fischer, P., Briggs, D., 2008. 430 A review of land-use regresssion models to assess spatial variation of outdoor air 431 pollution Atmospheric Environment 42, 7561 - 7578. 432 Hu, X., Belle, J.H., Meng, X., Wildani, A., Waller, L.A., Strickland, M.J., Liu, Y., 2017. 433 Estimating PM2.5 Concentrations in the Conterminous United States Using the 434 Random Forest Approach. Environ Sci Technol 51, 6936-6944. Kuhn, M., 2019. caret: Classification and Regression Training, R package version 6.0-84. 435 Kumar, A., Singh, D., Singh, B.P., Singh, M., Anandam, K., Kumar, K., Jain, V.K., 2015. 436 Spatial and temporal variability of surface ozone and nitrogen oxides in urban and 437 rural ambient air of Delhi-NCR, India. Air Quality, Atmosphere & Health 8, 391-399. 438 Li, H.Z., Gu, P., Ye, Q., Zimmerman, N., Robinson, E.S., Subramanian, R., Apte, J.S., 439 440 Robinson, A.L., Presto, A.A., 2019a. Spatially dense air pollutant sampling: 441 Implications of spatial variability on the representativeness of stationary air pollutant monitors. Atmospheric Environment: X 2, 100012. 442 Li, L., Girguis, M., Lurmann, F., Wu, J., Urman, R., Rappaport, E., Ritz, B., Franklin, M., 443 444 Breton, C., Gilliland, F., Habre, R., 2019b. Cluster-based bagging of constrained mixed-effects models for high spatiotemporal resolution nitrogen oxides prediction 445 446 over large regions. Environ Int 128, 310-323.

447 Lim, C.C., Kim, H., Vilcassim, M.J.R., Thurston, G.D., Gordon, T., Chen, L.C., Lee, K., 448 Heimbinder, M., Kim, S.Y., 2019. Mapping urban air quality using mobile sampling with low-cost sensors and machine learning in Seoul, South Korea. Environ Int 131, 449 450 105022. 451 Masiol, M., Squizzato, S., Chalupa, D., Rich, D.Q., Hopke, P.K., 2019. Spatial-temporal 452 variations of summertime ozone concentrations across a metropolitan area using a 453 network of low-cost monitors to develop 24 hourly land-use regression models. Sci 454 Total Environ 654, 1167-1178. 455 Masiol, M., Zikova, N., Chalupa, D.C., Rich, D.Q., Ferro, A.R., Hopke, P.K., 2018. Hourly 456 land-use regression models based on low-cost PM monitor data. Environ Res 167, 7-457 14. 458 Mead, M.I., Popoola, O.A.M., Stewart, G.B., Landshoff, P., Calleja, M., Hayes, M., Baldovi, 459 J.J., McLeod, M.W., Hodgson, T.F., Dicks, J., Lewis, A., Cohen, J., Baron, R., Saffell, 460 J.R., Jones, R.L., 2013. The use of electrochemical sensors for monitoring urban air 461 quality in low-cost, high-density networks. Atmospheric Environment 70, 186-203. Miskell, G., Alberti, K., Feenstra, B., Henshaw, G., Papapostolou, V., Patel, H., Polidori, A., 462 463 Salmond, J.A., Weissert, L.F., Williams, D.E., 2019. Reliable data from low-cost ozone sensors in a hierarchical network, http://arxiv.org/abs/1906.08421. 464 465 Miskell, G., Salmond, J., Alavi-Shoshtari, M., Bart, M., Ainslie, B., Grange, S., McKendry, 466 I.G., Henshaw, G.S., Williams, D.E., 2016. Data Verification Tools for Minimizing 467 Management Costs of Dense Air-Quality Monitoring Networks. Environ Sci Technol 468 50, 835-846. 469 Miskell, G., Salmond, J.A., Williams, D.E., 2018a. Solution to the Problem of Calibration of 470 Low-Cost Air Quality Measurement Sensors in Networks. ACS Sensors 3, 832-843.

Miskell, G., Salmond, J.A., Williams, D.E., 2018b. Use of a handheld low-cost sensor to

471

- 472 explore the effect of urban design features on local-scale spatial and temporal air 473 quality variability. Science of The Total Environment 619-620, 480-490. Popoola, O.A.M., Carruthers, D., Lad, C., Bright, V.B., Mead, M.I., Stettler, M.E.J., Saffell, 474 475 J.R., Jones, R.L., 2018. Use of networks of low cost air quality sensors to quantify air 476 quality in urban settings. Atmospheric Environment 194, 58-70. 477 Schneider, P., Castell, N., Vogt, M., Dauge, F.R., Lahoz, W.A., Bartonova, A., 2017. 478 Mapping urban air quality in near real-time using observations from low-cost sensors and model information. Environment International 106, 234-247. 479 480 Son, Y., Osornio-Vargas, Á.R., O'Neill, M.S., Hystad, P., Texcalac-Sangrador, J.L., Ohman-481 Strickland, P., Meng, Q., Schwander, S., 2018. Land use regression models to assess air pollution exposure in Mexico City using finer spatial and temporal input 482 483 parameters. Science of The Total Environment 639, 40-48. 484 South Coast AQMD, 2016. Final 2016 - Air Quality Management Plan. 485 Su, J.G., Jerrett, M., Beckerman, B., 2009. A distance-decay variable selection strategy for 486 land use regression modeling of ambient air pollution exposures. Sci Total Environ 487 407, 3890-3898. Vizcaino, P., Lavalle, C., 2018. Development of European NO2 Land Use Regression Model 488 489 for present and future exposure assessment: Implications for policy analysis. Environ 490 Pollut 240, 140-154. 491 Wei, Y., Wang, Y., Di, Q., Choirat, C., Wang, Y., Koutrakis, P., Zanobetti, A., Dominici, F., 492 Schwartz, J.D., 2019. Short term exposure to fine particulate matter and hospital 493 admission risks and costs in the medicare population: Time stratified, case crossover
- Weissert, L.F., Alberti, K., Miskell, G., Pattinson, W., Salmond, J.A., Henshaw, G.,

study. BMJ 367, 16258. 10.1136/bmj.16258

494

496	Williams, D.E., 2019a. Low-cost sensors and microscale land use regression: Data
497	fusion to resolve air quality variations with high spatial and temporal resolution.
498	Atmospheric Environment 213, 285-295.
499	Weissert, L.F., Miskell, G., Miles, E., Feenstra, B., Papapostolou, V., Polidori, A., Henshaw,
500	G.S., Salmond, J.A., Williams, D.E., 2019b. Hierarchical network design for
501	nitrogen dioxide measurement in urban environments, part 1: proxy selection.
502	(Available on: <a href="http://arxiv.org/abs/1911.03137">http://arxiv.org/abs/1911.03137</a> , Access Date: 4/12/19)
503	Weissert, L.F., Miles, E., Miskell, G., Alberti, K., Feenstra, B., Henshaw, G.S., Papapostolou,
504	V., Patel, H., Polidori, A., Polidori, A., Salmond, J.A., Williams, D.E, 2019c.
505	Hierarchical network design for nitrogen dioxide measurement in urban
506	environments, part 2: network-based sensor calibration. (Available on:
507	http://arxiv.org/abs/1911.03136, Access Date: 4/12/19)
508	WHO, 2013. Health risks of air pollution in Europe –HRAPIE project. Recommendations for
509	concentration-response functions for cost-benefit analysis of particulate matter,
510	ozone and nitrogen dioxide, World Health Organisation Regional Office for Europe,
511	Copenhagen
512	Williams, D.E., 2019. Low cost sensor networks: How do we know the data are reliable?
513	ACS Sensors 4, 2558-2565; <a href="https://doi.org/10.1021/acssensors.9b01455">https://doi.org/10.1021/acssensors.9b01455</a> .
514	Yeganeh, B., Hewson, M.G., Clifford, S., Tavassoli, A., Knibbs, L.D., Morawska, L., 2018.
515	Estimating the spatiotemporal variation of NO2 concentration using an adaptive
516	neuro-fuzzy inference system. Environmental Modelling & Software 100, 222-235.
517	Zhan, Y., Luo, Y., Deng, X., Zhang, K., Zhang, M., Grieneisen, M.L., Di, B., 2018. Satellite-
518	Based Estimates of Daily NO2 Exposure in China Using Hybrid Random Forest and
519	Spatio-temporal Kriging Model. Environ Sci Technol 52, 4180-4189.